

# A Time-Frequency Analysis of Oil Price Data

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October 3, 2017

# Purpose and summary

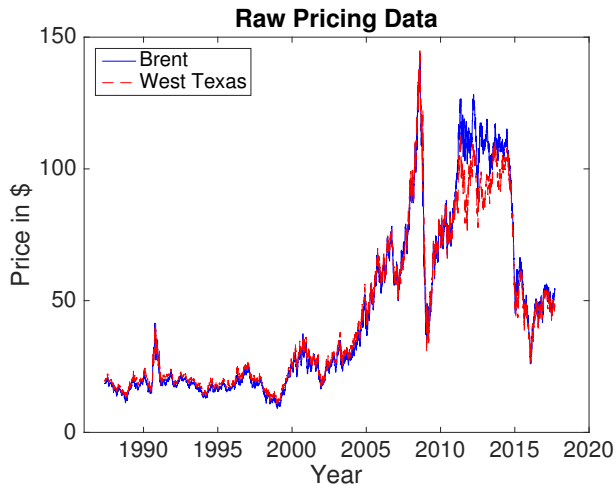
## Purpose:

- Challenge: to understand oil price dynamics in the long run, but also on short scales, in particular, for the recent period (since 2014).
- Idea: to adapt and exploit elaborated tools developed for turbulence data.

## Summary:

- Oil price data have a rich multiscale structure that may vary over time.
- The monitoring of these variations shows regime switches.
- The quantitative analysis is carried out by a wavelet decomposition method.

## The data set



Oil price data from 1987 to 2017 for West Texas (red dashed line) and Brent (solid blue line).

# Early history of price modeling

1900: **Louis Bachelier** “Théorie de la Spéculation”:

- Modeling of prices and pricing of options.
- Analysis of Brownian motion (predates Einstein’s 1905 works).

The price changes over different time intervals  $[t_n, t_{n+1}]$  are independent, Gaussian, zero-mean, and with variance proportional to  $t_{n+1} - t_n$ .

Increments of Brownian motion model “absolute” price changes, essentially:

$$P(t_{n+1}) - P(t_n) = \sigma(B(t_{n+1}) - B(t_n)).$$

# Standard price model

1960s: Paul A. Samuelson:

Increments of Brownian motion model “relative” price changes, essentially:

$$\frac{P(t_{n+1}) - P(t_n)}{P(t_n)} = \sigma(B(t_{n+1}) - B(t_n)),$$

with, in addition, possibly a deterministic drift.

## Standard price model

- Price  $P(t)$  = a drift  $d(t)$  + a diffusion, that can be expressed in terms of a Brownian motion  $B(t)$  and volatility  $\sigma_t$  :

$$\frac{dP(t)}{P(t)} = d(t)dt + \sigma_t dB(t)$$

- The Brownian motion  $B$  is a Gaussian process with **independent** and stationary increments:

$$\mathbb{E}[(B(t + \Delta t) - B(t))^2] = |\Delta t|$$

It is self-similar:

$$B(at) \stackrel{dist.}{\sim} a^{1/2}B(t) \quad \text{for any } a > 0$$

- The volatility is the most important ingredient of the standard model.

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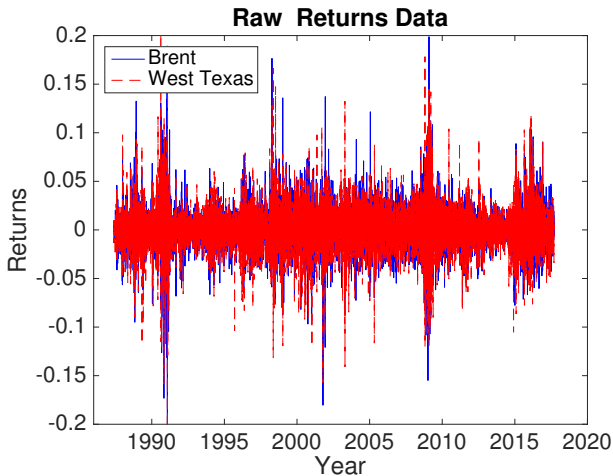
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- **The volatility is the most important ingredient of the standard model.**
  - ▶ When  $\sigma_t \equiv \sigma$  and  $d(t) \equiv d$ , this is the Black–Scholes model (1973).
  - ▶ When  $\sigma_t \equiv \sigma(t, P(t))$ , this is the local volatility model (Dupire, 1994; Derman and Kani, 1994).
  - ▶ When  $\sigma_t$  is a stochastic process, this is the stochastic volatility model (Hull and White, 1987; Heston, 1993).



Returns for West Texas (red dashed line) and Brent (solid blue line).

$$R_{n+1} = \frac{P(t_{n+1}) - P(t_n)}{P(t_n)}, \quad t_n = n\Delta t$$



# Modeling of prices: An alternative approach

1960s: **Benoit Mandelbrot**:

Increments of fractional Brownian motion model “relative” price changes, essentially:

$$R_{n+1} = \frac{P(t_{n+1}) - P(t_n)}{P(t_n)} = \sigma(B_H(t_{n+1}) - B_H(t_n)).$$

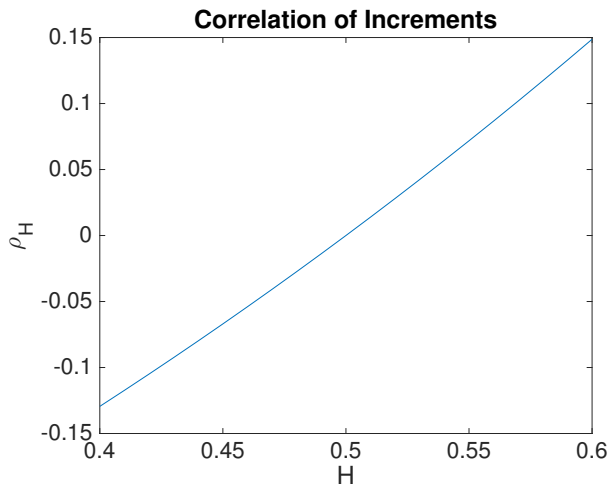
→ Returns have **“memory”**:

$$\rho_H = \frac{\mathbb{E}[R_{n+1}R_n]}{\mathbb{E}[R_n^2]} = 1 - 2^{2H-1}.$$

Thus in general:

$$\begin{aligned}\mathbb{E}[R_{n+1} \mid R_n] &= \rho_H R_n \\ &\neq \mathbb{E}[R_{n+1} \mid R_n, R_{n-1}].\end{aligned}$$

## Hurst sensitivity of return correlation



$$\rho_H = \frac{\mathbb{E}[R_{n+1}R_n]}{\mathbb{E}[R_n^2]} = 1 - 2^{2H-1}.$$

# Fractional price model

- Price  $P(t)$ :

$$\frac{dP(t)}{P(t)} = d(t)dt + \sigma_t dB_H(t)$$

where  $B_H$  is a fractional Brownian process with Hurst index  $H$ .

- $B_H$  is a Gaussian process with **dependent** and stationary increments:

$$\mathbb{E}[(B_H(t + \Delta t) - B_H(t))^2] = |\Delta t|^{2H}$$

- Properties:

- ▶ It is self-similar:

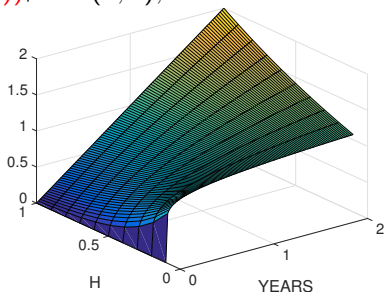
$$B_H(at) \stackrel{dist.}{\sim} a^H B_H(t) \quad \text{for any } a > 0$$

- ▶ If  $H = 1/2$ : it has uncorrelated increments (standard Brownian motion).
- ▶ If  $H < 1/2$ : it has negatively correlated increments (anti-persistence). Trajectories are rough (but continuous).
- ▶ If  $H > 1/2$ : it has positively correlated increments (persistence). Trajectories are smooth (but not differentiable).

Radical change of perspective: Hurst coefficient and volatility are the fundamental quantities.

Hurst coefficient governs scaling of volatility with time lag.

Example: we condition “volatility” to be one at time lag 1 (say in annualized units), then  $\text{St.dev}(\sigma B(t))$ ,  $t \in (0, 2)$ ,  $\sigma = 1$ :



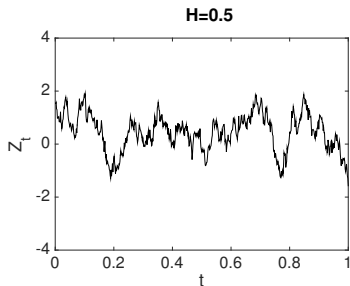
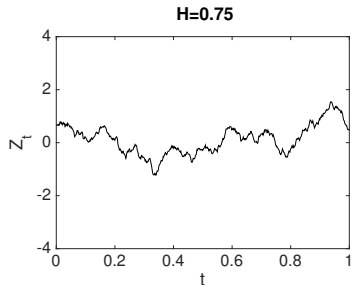
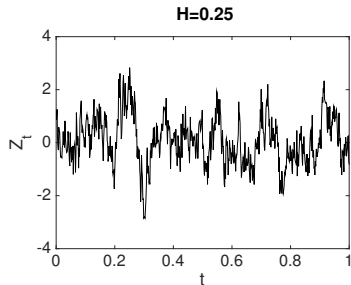
Classic case:

$H = 1/2$ : independent increments.

Limit cases:

$H \rightarrow 1$ : Increments equal:  $\Delta B_n = \Delta B_{n-1}$ .

$H \rightarrow 0$ : Increments alternate in sign:  $\Delta B_n = -\Delta B_{n-1}$ .



Ornstein-Uhlenbeck processes

$$dZ(t) = -Z(t)dt + dB_H(t)$$

# History

Fractional processes:

- Kolmogorov (1940): turbulence (turbulent flow composed by "eddies" of different sizes).
- Hurst (1951): hydrology (fluctuations of the water level in the Nile River).
- Mandelbrot and Van Ness (1968): finance.
- Comte and Renault (1998): stochastic volatility.
- Gatheral et al (2014): rough stochastic volatility.

# Multi-fractional price model

→ Motivated by the data, let increments of multi-fractional Brownian motion model “relative” price changes, essentially:

$$R_{n+1} = \frac{P(t_{n+1}) - P(t_n)}{P(t_n)} = \sigma_n(B_{H_n}(t_{n+1}) - B_{H_n}(t_n)).$$

- Price  $P(t)$ :

$$\frac{dP(t)}{P(t)} = d(t)dt + dB_{H,\sigma}(t)$$

where  $B_{H,\sigma}$  is a multi-fractional process ( $H_t$  and  $\sigma_t$  are time-dependent) (Lévy-Véhel 1995).

- If  $H_t \equiv H$  and  $\sigma_t \equiv 1$ , then  $B_{H,\sigma} = B_H$  fractional Brownian motion.

→ Some issues:

- Rapid Monte-Carlo simulation of price “paths”.
- Estimation of the parameters.  
↔ use of wavelets.

# Wavelets

- Main context: Signals may have frequency content that **varies with time**. Ex: speech.
- A **Fourier decomposition** gives the “global” frequency decomposition.
- A **wavelet decomposition** gives a local characterization of the frequency contents.

→ To detect changes in the multiscale character of the prices the wavelets are useful.

Cf: Meyer 1984, Mallat, Daubechies...



## Joint estimation of Hurst and volatility

- Use wavelet decomposition to localize in time and frequency the different components of the time series.
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  - ↔ The decay has a power-law form.
  - ↔ The parameters of this power law give the local Hurst and volatility parameters.

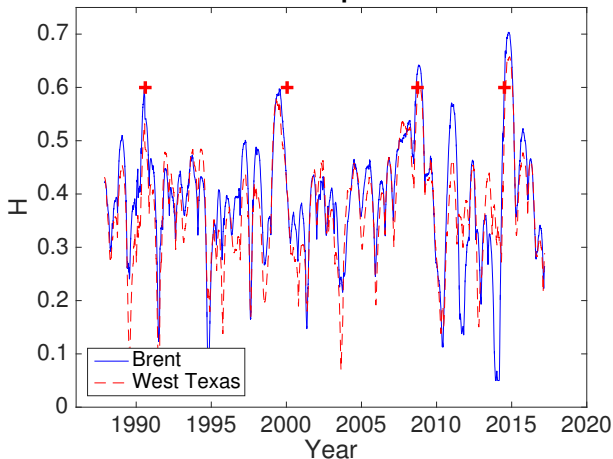
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- Remark: Many methods for Hurst parameter estimation:
  - ▶ Box-count estimator (Hall and Wood, 1993).
  - ▶ Variogram estimator (Constantine and Hall, 1994).
  - ▶ Level crossing estimator (Feuerverger, Hall and Wood, 1994).
  - ▶ Variation estimators (multiscale moments).
  - ▶ First spectral and wavelets estimators (Chan, Hall and Poskitt, 1995).

# Joint estimation of Hurst and volatility

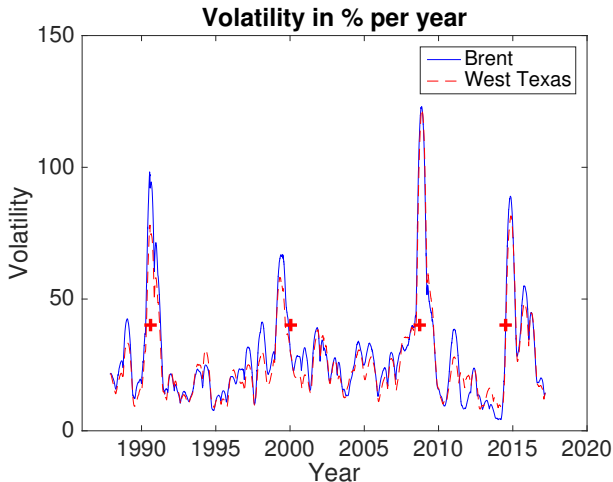
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  - ▶ First spectral and wavelets estimators (Chan, Hall and Poskitt, 1995).
- Statistical characterization of the local Hurst and volatility estimator.

## Hurst Exponent



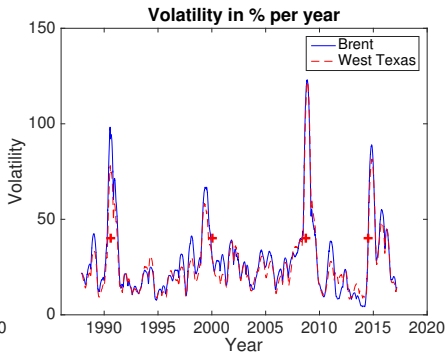
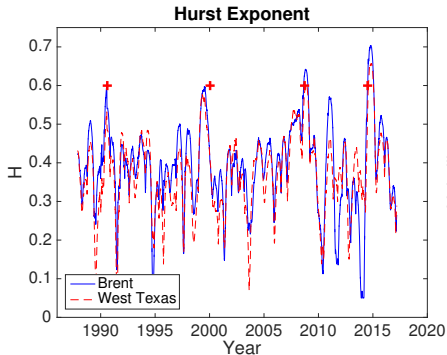
Estimated Hurst exponents  $H_t$ .

- There are four periods with a relatively high Hurst exponent, which can be related to four events, marked with crosses.

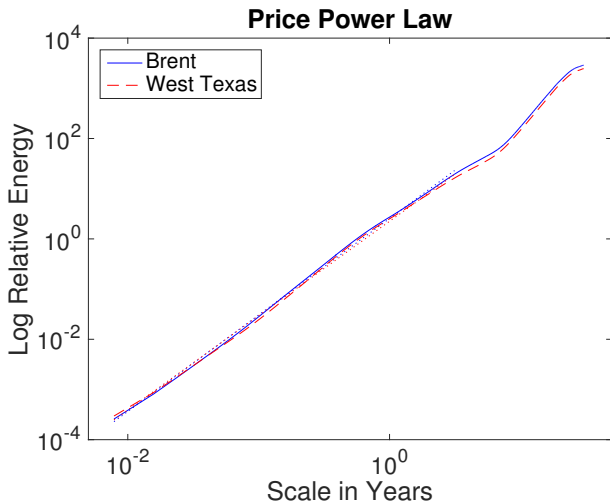


Local volatility estimates relative to the annual time scale  $\sigma_t$ .

- There are four periods with relatively high volatility, which can be related to four events, marked with crosses.
- Volatility is stable except: four special periods + period 2010-2014 (decaying volatility).



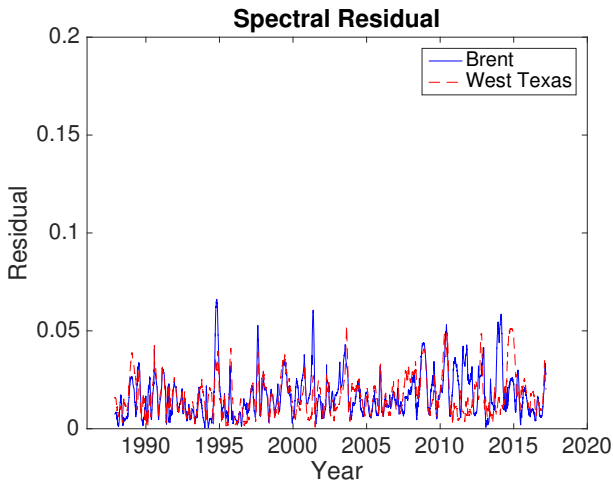
- **August 1990:** Iraq's invasion of Kuwait; it initiated a period with high volatility and a high Hurst exponent.
- **January 2000:** fear of the Y2K bug (?), which never occurred; it ended a period with relatively high volatility and Hurst exponent.
- **September 2008:** bankruptcy of Lehman Brothers; it initiated a period with very high volatility and a high Hurst exponent.
- **July 2014:** liquidation of oil-linked derivatives by fund managers; it initiated a period with a very high Hurst exponent and high volatility.



The “global power law” for West Texas data (red dashed line) and Brent data (blue solid line).

- A global power law (with  $H = .47$ ) is consistent with a situation in which the Hurst exponent and volatility vary over subsegments.





Spectral misfits for the West Texas data (red dashed line) and the Brent data (solid blue line).

- The spectral misfits are low and statistically homogeneous with respect to time.

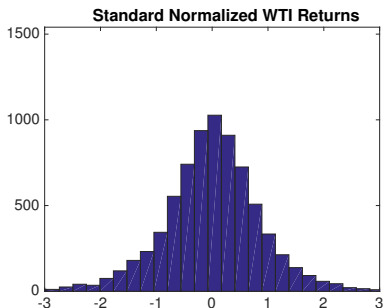
# Are returns Gaussian?

- Standard normalized returns:

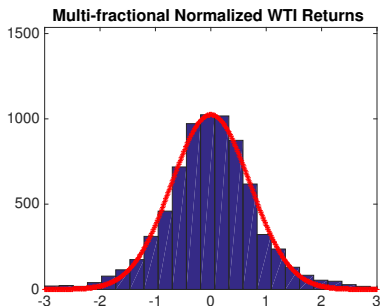
$$R_n^{(s)} = \frac{\log(P(t_n)) - \log(P(t_{n-1}))}{\sigma |\Delta t|^{1/2}}.$$

- Multi-fractional normalized returns:

$$R_n^{(m)} = \frac{\log(P(t_n)) - \log(P(t_{n-1}))}{\sigma_n |\Delta t|^{H_n}}.$$

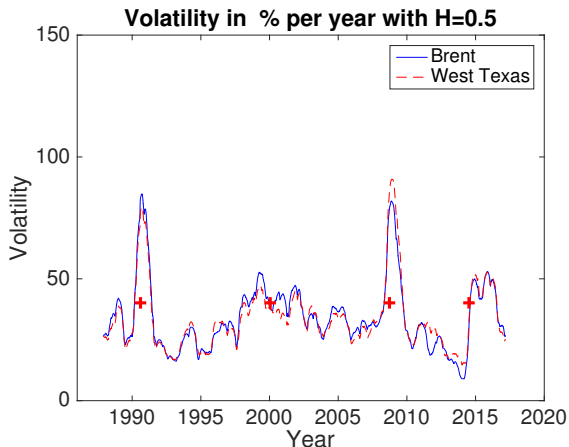


Standard normalized returns



Multi-fractional normalized returns

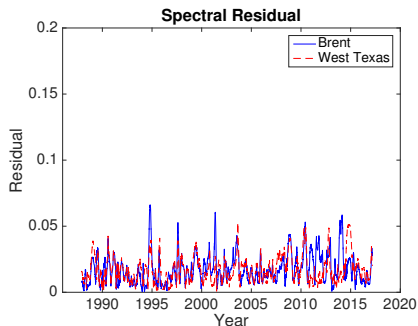
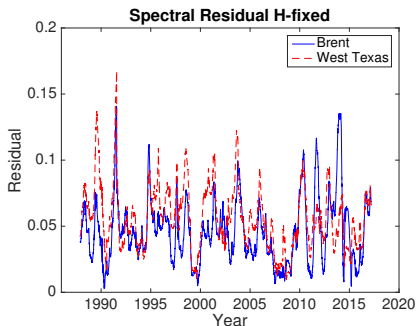
## Comparison with standard model



Estimated volatilities when we condition on  $H = 1/2$  to enforce uncorrelated returns.

- The four special periods do not appear so clearly; beyond these special periods, the standard volatility experiences somewhat strong variations.

# Comparison with standard model



Spectral misfits

**Left** when we condition on  $H = 1/2$  to enforce uncorrelated returns.

**Right** with the multi-fractional model.

- With  $H$  fixed, the spectral misfits are relatively high; they also vary significantly during the special periods.

## Conclusions so far

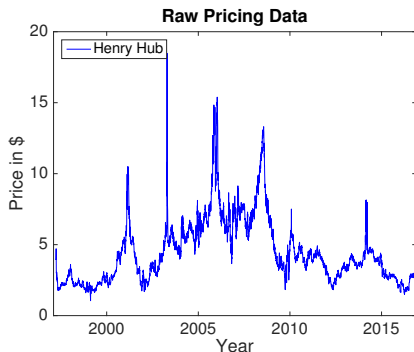
- Oil data contain multiscale fluctuations different from (log) normal diffusions.
- Special periods with  $H_t > 1/2$  can be identified.
- Standard volatility estimates are not appropriate during the special periods when  $H_t > 1/2$ .

## On classic and real markets

- In “classic” financial markets (with no arbitrage), the conditional expectation of the returns are zero.
- In “real” markets, we see deviations from the “ideal classical” context.
- For  $H \neq 1/2$ , the market is **not efficient** and pricing and hedging become challenging.

In what follows: different markets.

# Natural gas price

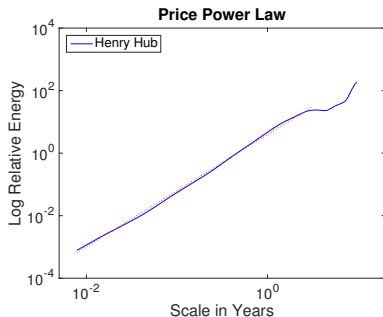
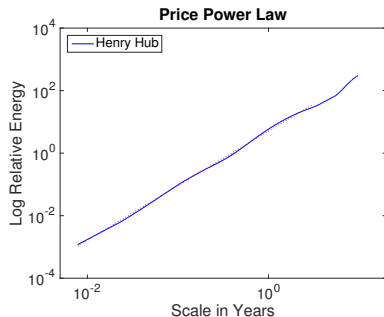


Henry Hub natural gas spot price.

## Some events:

- **2001**: California energy crisis;
- **2003**: cold winter;
- end of **2005**: hurricanes Katrina and Rita and volatile weather;
- **2008**: price increase corresponding to a high oil price.

# Natural gas price

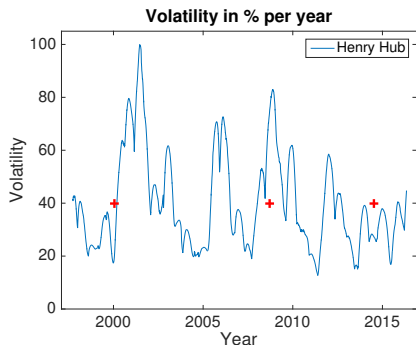
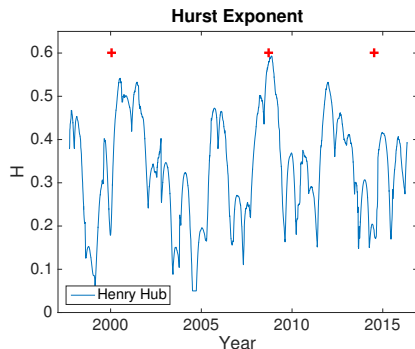


**Left:** Scale spectrum for Henry Hub natural gas for the period **1997–2009**. The estimated Hurst exponent is  $H = .37$ . The estimated volatility is  $\sigma = 49\%$ . The red dotted line is the fitted spectrum and fits well the data up to an outer scale of more than one year.

**Right:** Scale spectrum for Henry Hub natural gas for the period **2009–2016**. The estimated Hurst exponent is  $H = .40$ . The estimated volatility is  $\sigma = 43\%$ .



# Natural gas price



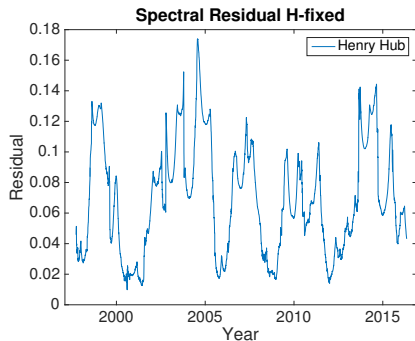
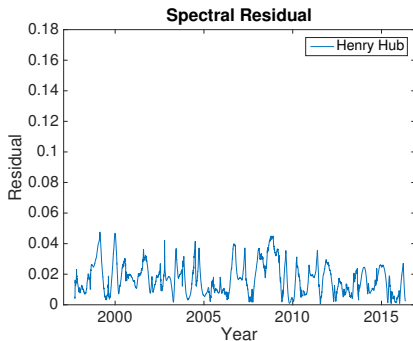
**Left:** Estimated Hurst exponents  $H_t$  for the Henry Hub natural gas spot price.

**Right:** Estimated volatility  $\sigma_t$  for the Henry Hub natural gas spot price.

Crosses:

**January 2000:** Y2K bug (?); **September 2008:** bankruptcy of Lehman Brothers; **July 2014:** liquidation of oil-linked derivatives by fund managers.

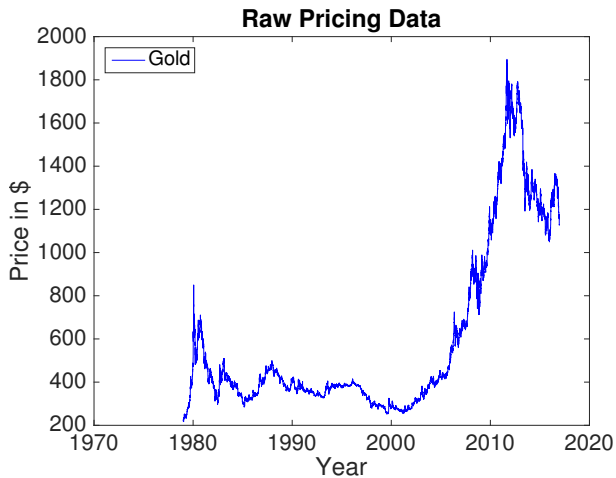
# Natural gas price



**Left:** Spectral misfits for Henry Hub natural gas.

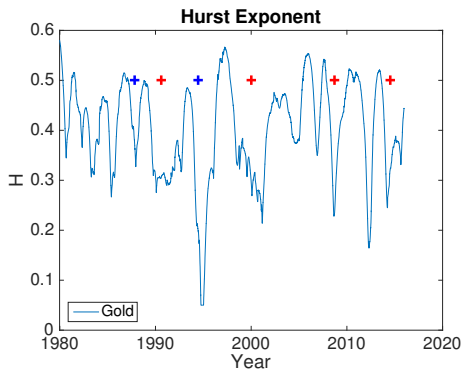
**Right:** Spectral misfits for Henry Hub natural gas when we condition on  $H = 1/2$  to enforce uncorrelated returns.

# Gold



Daily gold prices.

# Gold multi-fractal character



Daily gold prices.

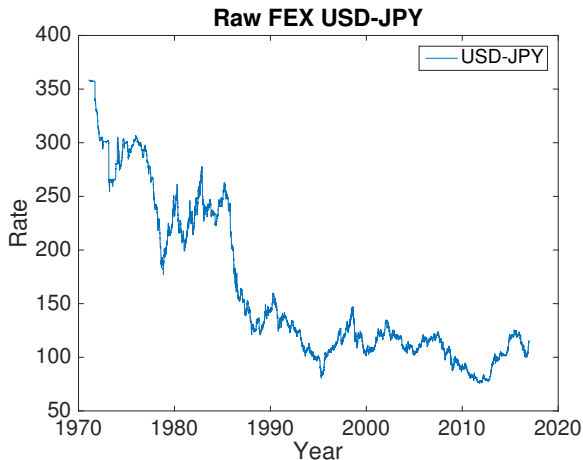
Blue crosses:

- Black Monday, October 19, 1987.
- Mexican peso crisis, December 1994.

Red crosses as above, but now associated with a rough epoch.

## Foreign exchange

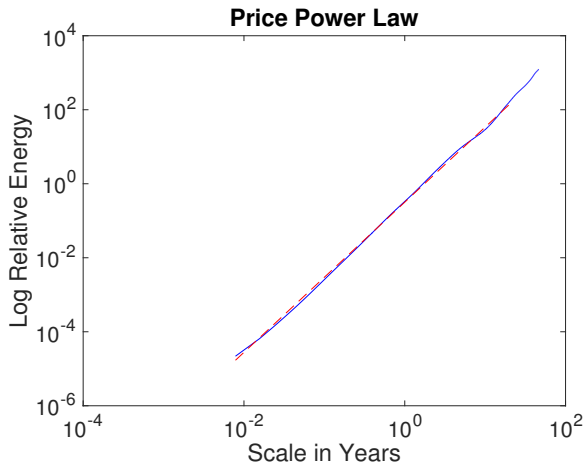
Consider the daily closing of yen-per-dollar price:



The price process corresponding to the daily closing of yen-per-dollar price.

- Does the exchange rate possess a multiscale structure?

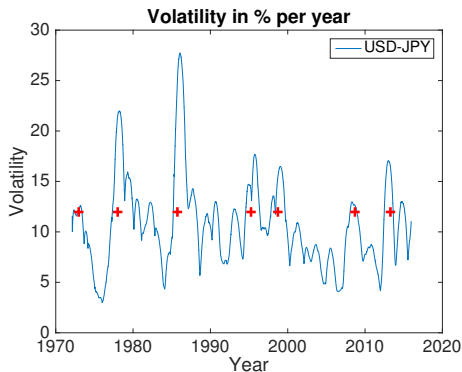
# USD–JPY scale spectrum



Scale spectrum for the full data set.

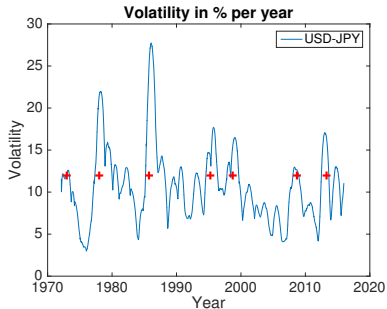
→ Efficiency on the grand scale with estimated  $H = .51$ .

# USD-JPY volatility



- **January 1973:** A series of events led to the first oil crisis that hit in October 1973;
- **January 1978:** A series of events led to the Iranian revolution of 1979 and the second oil crisis;
- **September 1985:** The *Plaza Accord*. This was a signed agreement between major nations affirming that the dollar was overvalued.

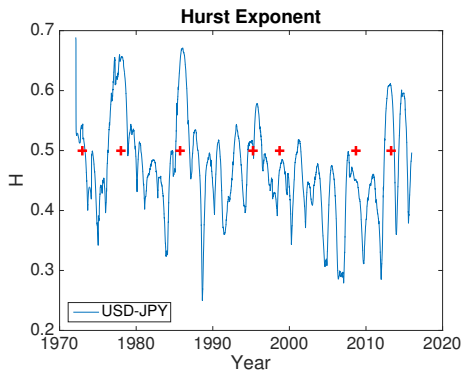
# USD–JPY volatility



- **April 1995:** The yen briefly hit a peak of under 80 yen per dollar after US–Japanese trade frictions sparked heavy selling of the dollar.
- **October 1998:** Near collapse of the hedge fund Long-Term Capital Management.
- **September 2008:** Lehman Brothers collapsed.
- **April 2013:** The Bank of Japan announced the expansion of its Asset Purchase program.

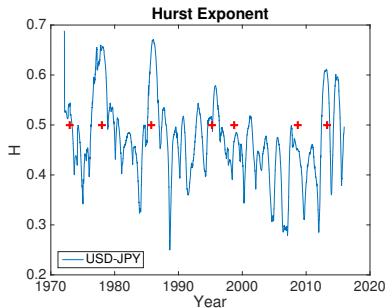


# USD-JPY Hurst



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## Some further issues

- “High-frequency” data.
- Periodicities.
- Cross-commodity correlations and derivatives.
- Interest rate.
- Consequences for speculation and risk management.

## Final remarks

- Multi-fractal behavior can be observed in various markets.
- We have developed a theory for the performance of the estimator.  
→ The Haar wavelets are partly superior.
- Multi-fractal modeling means a departure from classic financial modeling.
- Power-law modeling *potentially important* for regime shift detection, prediction, pricing, hedging,...
- Viewing the market and prices through the lens of both roughness and magnitude scaling  $(H, \sigma)$  gives “complementary” (economic) insight about the market.
- Claim: As in physics,  $H$  is a defining parameter.  $H < 1/2$  gives a modification of the efficient market situation, while  $H > 1/2$  changes the problem in a more fundamental way.

## Appendix: The case of dyadic Haar wavelets

Denote the **approximation coefficients** at level zero (the data) by:

$$\mathbf{X} = (a_0(1), a_0(2), \dots, a_0(2^M)).$$

Then, at the **scale**  $j$  (corresponding to frequency  $2^{-j}$ ), define the approximation and **difference coefficients** by:

$$a_j(n) = \frac{1}{\sqrt{2}}(a_{j-1}(2n) + a_{j-1}(2n-1))$$
$$d_j(n) = \frac{1}{\sqrt{2}}(a_{j-1}(2n) - a_{j-1}(2n-1)), \quad \text{for } n = 1, 2, \dots, 2^{M-j}$$

for  $j = 1, \dots, M$ .

## Appendix: Coefficients from the continuum

If  $a_0(n) = \int_{n-1}^n Y(t)dt$ , then the detail coefficients at level  $j$  can alternatively be expressed as:

$$d_j(n) = 2^{-j/2} \int_{-\infty}^{\infty} \psi(t2^{-j} - n)Y(t)dt$$

for  $Y$ , the (quasi-continuous) data, and with the mother wavelet defined by

$$\psi(x) = \begin{cases} -1 & \text{if } -1 \leq x < -1/2 \\ 1 & \text{if } -1/2 \leq x < 0 \\ 0 & \text{otherwise} \end{cases} .$$

→ The difference coefficients correspond to [probing the process at different scales and locations](#), with  $n$  representing the location and  $j$  the scale.

- From the self-similarity of fractional Brownian motion, it follows that for  $Y(t) = B_H(t)$ :

$$E[d_j(n)^2] \propto 2^{j(2H+1)}.$$

## Appendix: Scale spectrum

The scale spectrum of  $\mathbf{X}$ , relative to the Haar wavelet basis, is the sequence  $S_j$  defined by:

$$S_j = \frac{1}{2^{M-j}} \sum_{n=1}^{2^{M-j}} (d_j(n))^2, \quad j = 1, 2, \dots, M.$$

For fractional Brownian motion, the log of the scale spectrum is approximately *linear* in the scale  $j$ , with the slope determined by  $H$ .

## Appendix: The relation to the scale spectrum

- If the underlying process has the correlation structure of fractional Brownian motion, the log-scale spectrum is affine in scale:

$$\mathbb{E}[\log(S_j)] = c_0(\sigma, H) + (2H + 1)j.$$

Some issues:

- What are precision-of-parameter estimates?
- What is the role of the wavelets?
- What is the role of the scales and the shifts used?
- How should the regression be carried out?



## Appendix: Analysis of precision

- Different wavelets (Daubechies) can be used as well.
- A detailed analysis of the biases and variances of the volatility and Hurst parameter estimators is possible when the **underlying process is fractional Brownian motion**.
- The decomposition with Haar wavelets gives the most efficient estimator (as long as  $H$  is below .7).

## Appendix: On the parameter processes

Assume that the underlying process is fractional Brownian motion.

- It is possible to study the estimators  $\hat{H}(c)$  and  $\widehat{\log_2(\sigma^2)}(c)$  seen as processes indexed by the **right end** points  $c$  for moving time windows of equal length.
- The estimators  $\hat{H}(c)$  and  $\widehat{\log_2(\sigma^2)}(c)$ , as processes indexed by  $c$ , have **stationary Gaussian distributions** with a covariance structure that is **universal**.

→ It is possible to implement a filter to estimate more accurately the Hurst and volatility parameters.